# The Estimation of an Orientation Relationship 

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#### Abstract

A numerical method is given for the determination of an orientation relationship with high accuracy. It is essentially a least-squares method.


Conventional methods for the determination of an orientation relationship are usually carried out by means of manipulations on a stereographic net. However, the accuracy of these manipulations is essentially limited, and if high accuracy is required resort must be made, at least in the final stages, to purely numerical methods. In recent years the accuracy of orientation determinations has improved considerably, and in at least one investigation, concerned with a change of phase in the solid state (Bowles, Barrett \& Guttman, 1950), the accuracy approached 10 min . of arc. Thus, the need for purely numerical methods is apparent.

An orientation relationship is established when the rotation required to carry a standard orientation of a crystal into its actual orientation is known. If, on an actual crystal, exact measurements are made of the directions of at least two distinct normals with known specific indices this rotation can be determined. However, if the measurements are subject to error the angles between pairs of observed normals will not be exactly equal to their known true values and some adjustment procedure is necessary in order to estimate the rotation as precisely as possible. The method of adjustment proposed below was developed to solve a problem which arose in another connexion. It is numerical and analogous to the classical method of least squares for the solution of linear simultaneous equations. The method can be used whatever the precision of the observations but its main use lies in the final estimation of an orientation relationship of high accuracy.

The components of a vector relative to a fixed orthonormal basis can be written as the elements of a $3 \times 1$ column matrix. Therefore, let $y_{r}$ be the column matrix representing the measured unit normal to the $r$ th plane in the actual crystal and $\mathrm{x}_{r}$ that representing the known unit normal to the corresponding plane in the standard orientation. Then, if $R$ is the $3 \times 3$ rotation matrix ( $R R^{\prime}=1$ ) which is to be determined and $\theta_{r}\left(0 \leq \theta_{r} \leq \frac{1}{2} \pi\right)$ is the angular deviation of $y_{r}$ from its true position $\mathrm{Rx}_{r}$,

$$
\begin{equation*}
\cos \theta_{r}=y_{r}^{\prime} R x_{r}, \quad(r=1, \ldots, n) \tag{I}
\end{equation*}
$$

Now in the method of least squares $R$ would be deter-
mined so that a sum of the type $\Sigma w_{r} \theta_{r}^{2}$ was a minimum. However, for $\theta_{r}$ small, $\cos \theta_{r} \simeq 1-\frac{1}{2} \theta_{r}^{2}$ so that an almost equivalent procedure is to maximize the sum

$$
\begin{equation*}
S=\sum_{r=1}^{n} w_{r} \cos \theta_{r}=\sum_{r=1}^{n} w_{r} \mathrm{y}_{r}^{\prime} \mathrm{R} \mathrm{x}_{r} \tag{2}
\end{equation*}
$$

where each $w_{r}$ is a known (positive) weight which should ideally be chosen inversely proportional to the variance of $\theta_{r}$.

If $\mathrm{X}, \mathrm{Y}$ are $3 \times n$ matrices with $\mathrm{X}_{r}, \mathrm{y}_{r}$ in their $r$ th columns respectively and $W$ is the $n \times n$ diagonal matrix with diagonal elements $w_{r}$, then

$$
\begin{equation*}
S=\operatorname{Tr}\left(W Y^{\prime} R X\right)=\operatorname{Tr}(R A), \tag{3}
\end{equation*}
$$

where

$$
\begin{equation*}
A=X W Y^{\prime} \tag{4}
\end{equation*}
$$

and Tr denotes the sum of the diagonal elements of a matrix; the last step follows from the result that $\operatorname{Tr}(B C)=\operatorname{Tr}(C B)$ for all matrices $B, C$ which can be multiplied together. Now, for any matrix $A$, rotation matrices $R_{1}, R_{2}$ can be found such that

$$
\begin{equation*}
A=R_{2} \wedge R_{1}^{\prime} \tag{5}
\end{equation*}
$$

where $\Lambda$ is a diagonal matrix with non-negative elements and the product $R_{2} R_{1}^{\prime}$ is unique provided $A^{\prime} A$ has not more than one characteristic root equal to zero. Murnaghan (1938) effectively proves this result when $A$ is non-singular, but the general result, essentially proved below, is required since $A^{\prime} A$ will have just one zero characteristic root if measurements are made on only two normals. Thus,

$$
\begin{equation*}
S=\operatorname{Tr}\left(R_{1}^{\prime} R_{2} \Lambda\right)=\operatorname{Tr}\left(R_{0} \Lambda\right), \quad \text { say } \tag{6}
\end{equation*}
$$

Since the moduli of the elements of a rotation matrix are less than or equal to unity, $S$ is maximum when $R_{0}=I$ so that

$$
\begin{equation*}
R=R_{1} R_{2}^{\prime}=\left(R_{2} R_{1}^{\prime}\right)^{\prime} \tag{7}
\end{equation*}
$$

and

$$
\begin{equation*}
S_{\max .}=\operatorname{Tr}(\Lambda) \tag{8}
\end{equation*}
$$

## The procedure for determining R

The procedure is as follows: First the experimental data are used to compute $A$ from (4), then $R_{1}, R_{2}$ are
constructed as below and finally $R$ is given by (7). The final result is unique, apart from symmetry rotations.

The following construction of $R_{1}$ and $R_{2}$ is based on the facts that, if $A$ is given by (5), $R_{1}^{\prime} A^{\prime} A R_{1}=$ $R_{2}^{\prime} A A^{\prime} R_{2}=\Lambda^{2}$ and that $A^{\prime} A$ is symmetrical and has no negative characteristic roots, i.e. a positive semidefinite matrix (Albert, 1941). It follows (Ferrar, 1941, 1951) that a set of mutually orthogonal unit vectors $u_{1}, u_{2}, u_{3}$ can be determined such that

$$
\begin{equation*}
\mathrm{A}^{\prime} \mathrm{A} \mathrm{u}_{r}=\lambda_{r}^{2} \mathrm{u}_{r}, \quad \lambda_{r} \geq 0 \tag{9}
\end{equation*}
$$

where the $\lambda_{r}^{2}$ are the roots of the determinantal equation

$$
\begin{equation*}
\operatorname{det}\left[A^{\prime} A-\lambda^{2} I\right]=0 \tag{10}
\end{equation*}
$$

Thus, a possible choice for $R_{1}$ is

$$
\begin{equation*}
\mathrm{R}_{1}=\left(\mathrm{u}_{1}, \mathrm{u}_{2}, \mathrm{u}_{3}\right) \tag{11}
\end{equation*}
$$

and the diagonal elements of $\Lambda$ are $\lambda_{r}$. Multiplying (9) on the left by $A$ shows that, provided $\lambda_{r} \neq 0$,

$$
\begin{equation*}
\mathrm{v}_{r}=A \mathrm{u}_{r} \tag{12}
\end{equation*}
$$

is a characteristic vector of ${A A^{\prime}}^{\prime}$ belonging to the characteristic root $\lambda_{r}^{2}$. Further, it is easily shown that the $v_{r}$ so found are mutually orthogonal and of magnitude $\lambda_{r}$. Thus, when all the $\lambda_{r}$ are positive, a possible choice for $R_{2}$ is

$$
\begin{equation*}
\mathrm{R}_{2}=\left(\mathrm{v}_{1} / \lambda_{1}, \mathrm{v}_{2} / \lambda_{2}, \mathrm{v}_{3} / \lambda_{3}\right), \tag{13}
\end{equation*}
$$

and this is the only choice consistent with (5) and (11). If $\lambda_{3}=0$ is the only zero characteristic root, the first two columns of $R_{2}$ are found as above while the last column is the unique unit vector $v$ satisfying the equation

$$
\begin{equation*}
A^{\prime} \mathbf{v}=0 \tag{14}
\end{equation*}
$$

It can now be verified readily that $R_{1}^{\prime} A^{\prime} R_{2}=\Lambda$ so that $A$ is given by (5). Further, if all the characteristic
roots $\lambda_{r}$ are different the $u_{r}$ and $v_{r}$ are unique so that $R_{1}, R_{2}$ are each unique, while if the $\lambda_{r}$ are not all different it can be shown that the product $R_{1} R_{2}^{\prime}$ is unique although $R_{1}$ and $R_{2}$ separately are not.

The above estimate of $R$ can be made reasonable from the statistical point of view. For, if, following Fisher (1953), it is assumed that the observations are independent and that the probability density of each $\cos \theta_{r}$ is proportional to

$$
\begin{equation*}
\exp \left(\kappa w_{r} \cos \theta_{r}\right) \tag{15}
\end{equation*}
$$

where $\kappa w_{\tau}$ is approximately equal to the reciprocal of the variance of $\theta_{r}$, then the above estimate of R which maximizes $S$ is just the maximum likelihood estimate. Using this density function, the problems of determining the probability distributions of R and of $S_{\max }$. do not appear to be simple and no estimate of confidence limits can be made at present. However, the weighted mean of the expectation values of the $\cos \theta_{r}$ is approximately $1-1 / \mathrm{K} \operatorname{Tr}(W)$. Calling this quantity $\cos \theta_{0}$ and using the maximum likelihood estimate for $K$ gives

$$
\begin{equation*}
\cos \theta_{0}=1-\left[\operatorname{Tr}(W)-S_{\max .}\right] / n \operatorname{Tr}(W) \tag{16}
\end{equation*}
$$

and $\theta_{0}$ may be taken as a rough estimate of the 'reasonably likely' variations in the angle of rotation determined by $R$; $\theta_{0}$ will be related in some way to the standard error of this angle of rotation.

## References

Albert, A. A. (1941). Introduction to Algebraic Theories, p. 64. Chicago: University Press.

Bowles, J. S., Barrett, C. S. \& Guttman, L. (1950). Trans. Amer. Inst. Min. (Metall.) Engrs. 188, 1478.
Ferrar, W. L. (1941). Algebra, p. 131. Oxford: University Press.
Ferrar, W. L. (1951). Finite Matrices, p. 142. Oxford: University Press.
Fisher, R. A. (1953). Proc. Roy. Soc. A, 217, 295.
Murnaghan, F. D. (1938). Theory of Group Representations, p. 27. Baltimore: John Hopkins Press.

